

MATH180C: Introduction to Stochastic Processes II

Lecture A00: math.ucsd.edu/~ynemish/teaching/180cA

Lecture B00: math.ucsd.edu/~ynemish/teaching/180cB

Today: Martingales

Next: PK 8.1

Week 9:

- homework 7 (due Friday, May 27)

Maximal inequality for nonnegative martingales

Thm. Let $(X_n)_{n \geq 0}$ be a martingale with nonnegative values.

For any $\lambda > 0$ and $m \in \mathbb{N}$

$$P\left(\max_{0 \leq n \leq m} X_n \geq \lambda\right) \leq \frac{E(X_0)}{\lambda} \quad (1)$$

and

$$P\left(\max_{n \geq 0} X_n \geq \lambda\right) \leq \frac{E(X_0)}{\lambda} \quad (2)$$

Proof. We prove (1), (2) follows by taking the limit $m \rightarrow \infty$.

Take the vector (X_0, X_1, \dots, X_m) and partition the sample space wrt the index of the first r.v. rising above λ

$$I = \mathbb{1}_{X_0 \geq \lambda} + \mathbb{1}_{X_0 < \lambda, X_1 \geq \lambda} + \cdots + \mathbb{1}_{X_0 < \lambda, \dots, X_{m-1} < \lambda, X_m \geq \lambda} + \mathbb{1}_{X_0 < \lambda, \dots, X_m < \lambda}$$

Compute $E(X_m) = E(X_m \cdot I)$ using the above partition

Proof of the maximal inequality

$$\begin{aligned} E(X_m) &= \sum_{h=0}^m E(X_m \mathbb{1}_{X_0 < \lambda, \dots, X_{n-1} < \lambda, X_n \geq \lambda}) + E(X_m \mathbb{1}_{X_0 < \lambda, \dots, X_m < \lambda}) \\ &\geq \sum_{h=0}^m E(X_m \mathbb{1}_{X_0 < \lambda, \dots, X_{n-1} < \lambda, X_n \geq \lambda}) \end{aligned}$$

↗

Compute $E(X_m \mathbb{1}_{X_0 < \lambda, \dots, X_{n-1} < \lambda, X_n \geq \lambda})$ by conditioning on

$X_0, X_1, \dots, X_{n-1}, X_n$:

$$\begin{aligned} E(X_m \mathbb{1}_{X_0 < \lambda, \dots, X_{n-1} < \lambda, X_n \geq \lambda}) &= E(E(X_m \mathbb{1}_{X_0 < \lambda, \dots, X_{n-1} < \lambda, X_n \geq \lambda} | X_0, \dots, X_n)) \\ &= E(\mathbb{1}_{X_0 < \lambda, \dots, X_{n-1} < \lambda, X_n \geq \lambda} E(X_m | X_0, \dots, X_n)) \\ &= E(\mathbb{1}_{X_0 < \lambda, \dots, X_{n-1} < \lambda, X_n \geq \lambda} X_n) \geq \lambda \cdot P(X_0 < \lambda, \dots, X_{n-1} < \lambda, X_n \geq \lambda) \end{aligned}$$

Sum for all n

$$E(X_m) \geq \lambda \sum_{n=0}^m P(X_0 < \lambda, \dots, X_{n-1} < \lambda, X_n \geq \lambda) = \lambda P(\max_{0 \leq n \leq m} X_n \geq \lambda)$$

Example

A gambler begins with a unit amount of money and faces a series of independent fair games. In each game the gambler bets fraction p of his current fortune, wins with probability $\frac{1}{2}$, loses with probability $\frac{1}{2}$.

Estimate the probability that the gambler ever doubles the initial fortune.

Denote by $Z_n, n \geq 0$, the gambler's fortune after n -th game.

Denote $\{Y_i\}_{i=1}^{\infty}$ i.i.d. r.v.s with $P(Y_i = 1+p) = P(Y_i = 1-p) = \frac{1}{2}$

Then $Z_n = Y_1 \cdot Y_2 \cdots \cdot Y_n, Z_0 = 1$

$E(Y_i) = (1+p) \cdot \frac{1}{2} + (1-p) \cdot \frac{1}{2} = 1 \Rightarrow (Z_n)_{n \geq 0}$ is a nonnegative martingale

$$\Rightarrow P\left(\max_{n \geq 0} Z_n \geq 2\right) \leq \frac{E(Z_0)}{2} = \frac{1}{2}$$

Martingale transform

In the previous example the stake in n-th game is

$P Z_{n-1}$. What if we choose another strategy?

Def Let $(X_n)_{n \geq 0}$ be a nonnegative martingale, and let $(C_n)_{n \geq 0}$ be a stochastic process with

$C_n = f_n(X_0, \dots, X_{n-1})$. Then the stochastic process

$$\sum_{k=1}^n C_k (X_k - X_{k-1}) = (C \bullet X)_n, \quad (C \bullet X)_0 = 0$$

is called the martingale transform of X by C

Think of • $X_k - X_{k-1}$ as the winning per unit stake in k-th game

• C_k as your stake in k-th game

decision is made based on the previous history

• $(C \bullet X)_n$ as total winnings up to time n

Martingale transform

Prop. Let $Z_n = X_0 + (C \cdot X)_n$. Let $C_k > 0$ bounded if $Z_{k-1} > 0$ and $C_k = 0$ if $Z_{k-1} = 0$. Then $(Z_n)_{n \geq 0}$ is a martingale

Proof: $E(Z_{n+1} | Z_0, \dots, Z_n) = E(Z_n + C_{n+1}(X_{n+1} - X_n) | Z_0, \dots, Z_n)$
 $= Z_n + E(C_{n+1}(X_{n+1} - X_n) | Z_0, \dots, Z_n)$

Note that $Z_n - Z_{n-1} = C_n(X_n - X_{n-1})$, $Z_0 = X_0$

If $Z_n > 0$, then $C_1 > 0, \dots, C_n > 0$,

$$X_1 = (Z_1 - Z_0)C_1^{-1} + Z_0, \quad X_n = (Z_n - Z_{n-1})C_n^{-1} + X_{n-1} \text{ and}$$

$$\begin{aligned} E(Z_{n+1} | Z_0, \dots, Z_n) &= Z_n + E(C_{n+1}(X_{n+1} - X_n) | X_0, \dots, X_n) \\ &= Z_n \end{aligned}$$

If $Z_n = 0$, then $C_{n+1} = 0$ and $E(Z_{n+1} | Z_0, \dots, Z_n) = 0 = Z_n$



Gambling example:

Start from the initial fortune $X_0 = 1$. Define

$$Z_n = 1 + (C \bullet X)_n$$

fortune after n -th game with strategy C

Then $(Z_n)_{n \geq 0}$ is a nonnegative martingale, $E(Z_0) = 1$

$$\Rightarrow P\left(\max_{n \geq 0} Z_n \geq 2\right) \leq \frac{1}{2}$$

Convergence of nonnegative martingales

Thm.

If $(X_n)_{n \geq 0}$ is a nonnegative (super)martingale, then with probability 1

$$\exists \lim_{n \rightarrow \infty} X_n =: X_\infty$$

and

$$E(X_\infty) \leq E(X_0)$$

Example

An urn initially contains one red ball and one green ball. Choose a ball and return it to the urn together with another ball of the same color. Repeat. Denote by X_n the fraction of red ball after n iterations.

Example (cont.)

(i) $(X_n)_{n \geq 0}$ is a martingale

Denote by R_n the number of red balls after n -th iteration

$$R_n = X_n \cdot (n+2)$$

Then

$$\begin{aligned} E(X_{n+1} | X_0, \dots, X_n) &= \frac{R_n + 1}{n+3} X_n + \frac{R_n}{n+3} (1 - X_n) \\ &= \frac{1}{n+3} (X_n + R_n) = \frac{1}{n+3} (X_n + X_n(n+2)) = X_n \end{aligned}$$

(ii) X_n is nonnegative $\Rightarrow \exists \lim_{n \rightarrow \infty} X_n =: X_\infty$

(iii) Compute the distribution of X_∞

$$P(X_n = \frac{k}{n+2}) = \frac{1}{n+1} \quad \text{for } k \in \{1, 2, \dots, n+1\}$$

$$P(X_\infty \leq x, x \in (0,1)) \Rightarrow X_\infty \sim \text{Unif}[0,1]$$