

# MATH180C: Introduction to Stochastic Processes II

[Lecture A00: math-old.ucsd.edu/~ynemish/teaching/180cA](http://math-old.ucsd.edu/~ynemish/teaching/180cA)

[Lecture B00: math-old.ucsd.edu/~ynemish/teaching/180cB](http://math-old.ucsd.edu/~ynemish/teaching/180cB)

Today: Poisson process as a  
renewal process. Other  
examples

Next: PK 7.4-7.5, Durrett 3.1

Week 6:

- homework 5 (due Friday, May 6)
- regrades for Midterm 1 active until May 7, 11PM

## Renewal density

$$f * f(t) = \int f(t-x)f(x) dx$$

Proposition Let  $N(t)$  be a renewal process with continuous interrenewal times  $X_i$  having density  $f(x)$ . Denote

$$m(t) = \sum_{n=1}^{\infty} f^{*n}(t) \quad \text{Then} \quad M(t) = \int_0^t m(x) dx$$

$$\text{and} \quad m(t) = f(t) + m * f(t) \quad (*)$$

↑ renewal density

Proof:  $\frac{d}{dt} F^{*n}(t) = \left( \frac{d}{dt} F^{*(n-1)} \right) * f(t) = f^{*n}(t) \quad \blacksquare$

Example: Compute the renewal density for PP using (\*).

$f(x) = \lambda e^{-\lambda x}$ , so (\*) becomes

$$\begin{aligned} m(t) &= \lambda e^{-\lambda t} + \int_0^t m(t-x) \lambda e^{-\lambda x} dx = \lambda e^{-\lambda t} + \int_0^t m(x) \lambda e^{-\lambda(t-x)} dx \\ &= \lambda e^{-\lambda t} \left( 1 + \int_0^t m(x) e^{\lambda x} dx \right) \end{aligned}$$

(cont.)

$$e^{\lambda t} m(t) = \lambda \left( 1 + \int_0^t m(x) e^{\lambda x} dx \right) \leftarrow \text{differentiate}$$

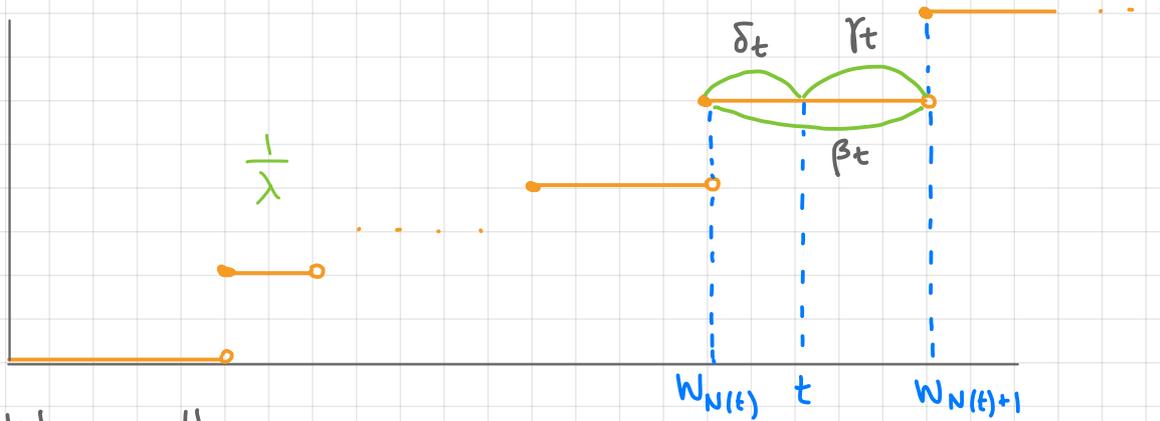
$$\left\{ \begin{array}{l} \frac{d}{dt} (e^{\lambda t} m(t)) = \lambda (m(t) e^{\lambda t}) \\ m(0) = \lambda \end{array} \right. \Rightarrow e^{\lambda t} m(t) = \lambda e^{\lambda t}$$

$$m(t) = \lambda$$

Indeed, 
$$M(t) = \int_0^t m(x) dx = \int_0^t \lambda dx = \lambda t$$

# Excess life and current life of PP (summary)

Recall: Let  $N(t)$  be a renewal process.



Def. We call

- $\gamma_t := W_{N(t)+1} - t$  the excess (or residual) lifetime
- $\delta_t := t - W_{N(t)}$  the current life (or age)
- $\beta_t := \gamma_t + \delta_t$  the total life

Remarks 1)  $\gamma_t > h \geq 0$  iff  $N(t+h) = N(t)$   
2)  $t \geq h$  and  $\delta_t \geq h$  iff  $N(t-h) = N(t)$

## Excess life and current life of PP

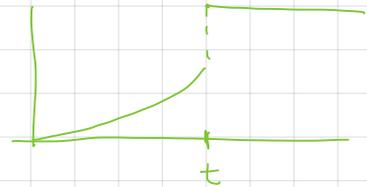
Let  $N(t)$  be a PP. Then

- excess life  $\gamma_t \sim \text{Exp}(\lambda)$

$$P(\gamma_t > x) = P(N(t+x) - N(t) = 0) = P(N(x) = 0) = e^{-\lambda x}$$

- current life  $\delta_t$

$$P(\delta_t > x) = \begin{cases} 0, & x \geq t \\ P(N(t-x) = N(t)) = P(N(t) - N(t-x) = 0) = e^{-\lambda x}, & x < t \end{cases}$$



- total life  $\beta_t = \gamma_t + \delta_t$

$$\begin{aligned} E(\gamma_t + \delta_t) &= \frac{1}{\lambda} + E(\delta_t) = \frac{1}{\lambda} + \int_0^{\infty} P(\delta_t > x) dx \\ &= \frac{1}{\lambda} + \int_0^t e^{-\lambda x} dx = \frac{1}{\lambda} + \frac{1}{\lambda} (1 - e^{-\lambda t}) \rightarrow \frac{2}{\lambda} \text{ as } t \rightarrow \infty \end{aligned}$$

## Excess life and current life of PP (cont.)

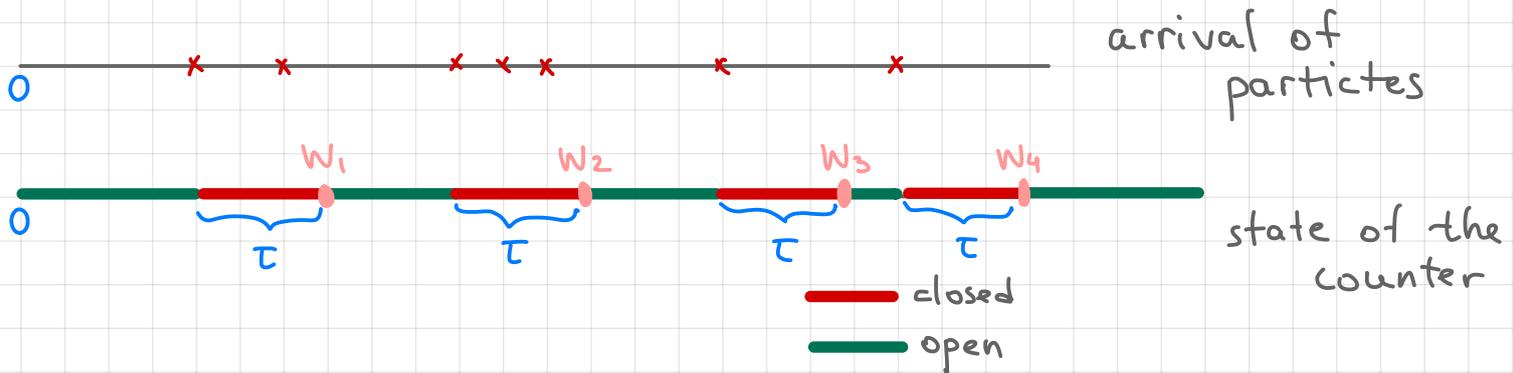
- Joint distribution of  $(\gamma_t, \delta_t)$

$$P(\gamma_t > x, \delta_t > y) = \begin{cases} 0, & \text{if } y \geq t \\ P(N(t-y) = N(t+x)) = e^{-\lambda(x+y)}, & \text{if } y < t \end{cases}$$

$\Rightarrow \gamma_t$  and  $\delta_t$  are independent (for PP)

## Other renewal processes

- traffic flow : distances between successive cars are assumed to be i.i.d. random variables
- counter process: particles/signals arrive on a device and lock it for time  $\tau$  ; particles arrive according to a PP; times at which the counter unlocks form a renewal process



## Other renewal processes

- more generally, if a component has two states (0/1, operating/non-operating etc), switches between them, times spent in 0 are  $X_i$ , times spent in 1 are  $Y_i$ ,  $(X_i)_{i=1}^{\infty}$  i.i.d.,  $(Y_i)_{i=1}^{\infty}$  i.i.d., then the times of switching from 0 to 1 form a renewal process with interrenewal times  $X_i + Y_i$



## Other renewal processes

- Markov chains: if  $(Y_n)_{n \geq 0}$ ,  $Y_n \in \{0, 1, \dots\}$  is a recurrent MC starting from  $Y_0 = k$ , then the times of returns to state  $k$  form a renewal process. More precisely

$$\text{define } W_1 = \min \{n > 0 : Y_n = k\}$$

$$W_p = \min \{n > W_{p-1} : Y_n = k\}$$



Example with  $k=2$

Similarly for continuous time MCs.

Strong Markov property!

## Other renewal processes

- Queues. Consider a single-server queueing process



- (i) if customer arrival times form a renewal process then the times of the starts of successive idle periods generate a second renewal time
- (ii) if customers arrive according to a Poisson process, then the times when the server passes from busy to free form a renewal process