

MATH 285: Stochastic Processes

math-old.ucsd.edu/~ynemish/teaching/285

Today: Martingales

- Homework 6 is due on Friday, March 4, 11:59 PM

Martingales

Def 24.1 A discrete-time martingale is a stochastic process $(X_n)_{n \geq 0}$ which satisfies $\mathbb{E}[|X_n|] < \infty$ and

$$\mathbb{E}[X_{n+1} | X_0, \dots, X_n] = X_n \text{ for all } n \geq 0$$

Lemma 24.2 If $(X_n)_{n \geq 0}$ is a martingale, then

$$\mathbb{E}[X_n | X_0, \dots, X_m] = X_m \text{ for all } m < n.$$

Proof. Fix m . Induction. Holds for $n=m$, $n=m+1$.

Suppose $\mathbb{E}[X_n | X_0, \dots, X_m] = X_m$ for some $n > m$. Then by the Tower property

$$\begin{aligned}\mathbb{E}[X_{n+1} | X_0, \dots, X_m] &= \mathbb{E}[\mathbb{E}[X_{n+1} | X_0, \dots, X_n] | X_0, \dots, X_m] \\ &= \mathbb{E}[X_n | X_0, \dots, X_m] = X_m\end{aligned}$$

■

Martingales

Corollary 24.3 If $(X_n)_{n \geq 0}$ is a martingale, then it has constant expectation: $\mathbb{E}[X_n] = \mathbb{E}[X_0]$ for all n .

Proof. Use the double expectation property

$$\mathbb{E}[X_n] = \mathbb{E}[\mathbb{E}[X_n | X_0]] = \mathbb{E}[X_0] \quad \blacksquare$$

Example (Betting on independent coin tosses)

Consider a game: bet B_i dollars and toss a coin.

$$X_i = \begin{cases} 1, & \text{if you win the } i\text{-th toss} \\ -1, & \text{if you lose the } i\text{-th toss} \end{cases}, \quad X_1, X_2, \dots \text{ independent}$$

Denote by W_0 the initial fortune, independent of X_1, X_2, \dots

Let $W_n = W_0 + \sum_{i=1}^n X_i B_i$. We call B_1, B_2, \dots the betting strategy.

- $\mathbb{E}[|B_n|] < \infty$
- B_n is (W_0, \dots, W_{n-1}) -measurable

Betting on independent coin tosses

Then

$$\mathbb{E}[|W_n|] \leq \mathbb{E}[|W_0| + \sum_{i=1}^n |X_i B_i|] = \mathbb{E}[|W_0|] + \sum_{i=1}^n \mathbb{E}[|B_i|] < \infty \quad \forall n$$

and

$$\begin{aligned}\mathbb{E}[W_{n+1} | W_0, \dots, W_n] &= \mathbb{E}[W_n + B_{n+1} X_{n+1} | W_0, \dots, W_n] \\ &= W_n + \mathbb{E}[X_{n+1} B_{n+1} | W_0, \dots, W_n] \\ &= W_n + B_{n+1} \mathbb{E}[X_{n+1} | W_0, \dots, W_n]\end{aligned}$$

Since $W_K = \sum_{i=1}^K X_i B_i$, X_{n+1} is independent of W_0, \dots, W_n

||| W_0 is independent of X_{n+1}

B_1 is W_0 measurable, $W_1 = W_0 + X_1 B_1$

B_2 is (W_0, W_1) measurable, $W_2 = W_1 + X_2 B_2 \dots$

Then $\mathbb{E}[X_{n+1} | W_0, \dots, W_n] = \mathbb{E}[X_{n+1}] = 0$, and $\mathbb{E}[W_{n+1} | W_0, \dots, W_n] = W_n$
 i.e., W_n is a martingale

Stopping times

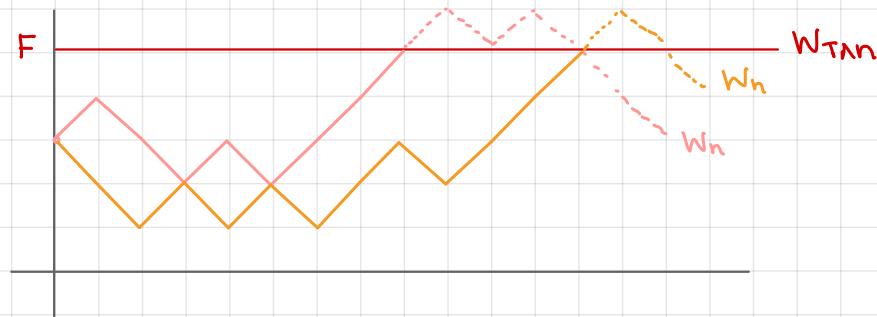
Let $(W_n)_{n \geq 0}$ be a stochastic process. Recall that random variable $T \in \{0, 1, 2, \dots\} \cup \{\infty\}$ is a stopping time if the fact that $\{T \leq n\}$ holds can be determined from W_0, \dots, W_n .

Example of a stopping times: the first time the process hits some set/value, $T = \min\{n \geq 0 : W_n \geq F\}$

Suppose you stop the game as soon as your fortune gets $\geq F$. Then the original process (W_n) is replaced by $(W_{T \wedge n})_{n \geq 0}$,

where $T \wedge n = \min\{T, n\}$

$$W_{T \wedge n} = \begin{cases} W_n, & \text{if } n < T \\ W_T, & \text{if } n \geq T \end{cases}$$



Stopped martingale

Prop. 24.5 Let $(X_n)_{n \geq 0}$ be a martingale and let T be a stopping time for this martingale. Then

$(X_{T \wedge n})_{n \geq 0}$ is a martingale.

Proof. Denote $Y_n := X_{T \wedge n}$. Then $Y_n = X_{T \wedge n} \in \{X_0, X_1, \dots, X_n\}$ for each n , so $|Y_n| \leq \max\{|X_0|, \dots, |X_n|\} \leq |X_0| + \dots + |X_n|$, and

$$\mathbb{E}[|Y_n|] \leq \mathbb{E}[|X_0| + \dots + |X_n|] < \infty$$

Now we need to show that $\mathbb{E}[Y_{n+1} | Y_0, \dots, Y_n] = Y_n$.

(1) $\mathbb{E}[Y_{n+1} | X_0, \dots, X_n] = Y_n$

- $Y_{n+1} = X_{T \wedge (n+1)} = X_T \mathbf{1}_{\{T \leq n\}} + X_{n+1} \mathbf{1}_{\{T > n\}}$

- $\{T \leq n\}$ only depends on X_0, \dots, X_n , so $\mathbf{1}_{\{T \leq n\}}$ is (X_0, \dots, X_n) -measurable and $X_T \mathbf{1}_{\{T \leq n\}}$ is (X_0, \dots, X_n) measurable

Stopped martingale

- $\mathbb{1}_{\{T>n\}} = 1 - \mathbb{1}_{\{T \leq n\}}$ is (X_0, \dots, X_n) -measurable
- Using the properties of the conditional expectation

$$\begin{aligned}\mathbb{E}[Y_{n+1} | X_0, \dots, X_n] &= \mathbb{E}[X_T \mathbb{1}_{\{T \leq n\}} | X_0, \dots, X_n] + \mathbb{E}[X_{n+1} \mathbb{1}_{\{T>n\}} | X_0, \dots, X_n] \\ &= X_T \mathbb{1}_{\{T \leq n\}} + \mathbb{1}_{\{T>n\}} \mathbb{E}[X_{n+1} | X_0, \dots, X_n] \\ &= X_T \mathbb{1}_{\{T \leq n\}} + \mathbb{1}_{\{T>n\}} X_n\end{aligned}$$

- $X_T \mathbb{1}_{\{T \leq n\}} + X_n \mathbb{1}_{\{T>n\}} = X_T \mathbb{1}_{\{T \leq n\}} + \mathbb{1}_{\{T \geq n\}} X_n = Y_n$

(2) $\mathbb{E}[Y_{n+1} | Y_0, \dots, Y_n] = Y_n$

- $\bar{Y}_n := (Y_0, \dots, Y_n)$ is $\bar{X}_n := (X_0, \dots, X_n)$ measurable

- Using the Tower property $\mathbb{E}[Y_{n+1} | \bar{Y}_n] = \mathbb{E}[\mathbb{E}[Y_{n+1} | \bar{X}_n] | \bar{Y}_n]$

- By (1) $\mathbb{E}[\mathbb{E}[Y_{n+1} | \bar{X}_n] | \bar{Y}_n] \stackrel{(1)}{=} \mathbb{E}[Y_n | \bar{Y}_n] = Y_n$

Martingale betting strategy

Corollary For any $n \in \mathbb{N}$ $\mathbb{E}[X_{T \wedge n}] = \mathbb{E}[X_{T \wedge 0}] = \mathbb{E}[X_0]$

Example (Martingale betting strategy)

$$\mathbb{E}[X_T]$$

Consider the betting strategy $B_n = 2^{n-1}$ (double each round).

Let $T = \min\{n : X_n = 1\}$, the time of the first win. If $W_0 = C$,

then

$$\mathbb{E}[W_T] = \sum_{n=1}^{\infty} \mathbb{E}[W_T | T=n] \mathbb{P}[T=n]$$

The event $T=n$ corresponds to a specific trajectory

$$X_1 = -1, X_2 = -1, \dots, X_{n-1} = -1, X_n = 1, \text{ so } \mathbb{P}[T=n] = \frac{1}{2^n}$$

$$\begin{aligned} \text{So } \mathbb{E}[W_T] &= \sum_{n=1}^{\infty} \mathbb{E}[W_n | X_1 = -1, \dots, X_{n-1} = -1, X_n = 1] \frac{1}{2^n} \\ &= \sum_{n=1}^{\infty} \left(C - 1 - 2 - 4 - \dots - 2^{n-2} + 2^{n-1} \right) \frac{1}{2^n} = C + 1 \neq C = \mathbb{E}[W_0] \end{aligned}$$

Problem T can be arbitrarily large.

Optional Sampling Theorem

Thm 24.8 Let $(X_n)_{n \geq 0}$ be a martingale, and let T be a finite stopping time. Suppose that either

(1) T is bounded : $\exists N < \infty$ s.t. $P[T < N] = 1$; or

(2) $(X_n)_{0 \leq n \leq T}$ is bounded : $\exists B < \infty$ s.t. $P[|X_n| \leq B \text{ for all } n \leq T] = 1$

Then $E[X_T] = E[X_0]$

Proof. • Suppose (1) holds. By Prop. 24.5 $X_{T \wedge n}$ is a martingale, $E[X_{T \wedge n}] = E[X_{T \wedge 0}] = E[X_0]$ for all n .

Then $E[X_0] = E[X_{T \wedge N}] = E[X_T]$

• Suppose that (2) holds (T is not necessarily bounded)

Then $X_T = X_{T \wedge n} + (X_T - X_{T \wedge n}) = X_{T \wedge n} + (X_T - X_{T \wedge n}) \mathbb{1}_{\{T > n\}}$

First term: $E[X_{T \wedge n}] = E[X_0]$

Optional Sampling Theorem

Second term:

$$\begin{aligned} |\mathbb{E}[(X_T - X_{T \wedge n}) \mathbb{1}_{\{T > n\}}]| &\leq \mathbb{E}[|X_T - X_{T \wedge n}| \mathbb{1}_{\{T > n\}}] \\ &\leq 2B \mathbb{P}[T > n] \end{aligned}$$

Since $\mathbb{P}[T < \infty] = 1$, $\lim_{n \rightarrow \infty} \mathbb{P}[T > n] = 0$

Therefore,

$$\begin{aligned} |\mathbb{E}[X_T] - \mathbb{E}[X_0]| &= |\mathbb{E}[(X_T - X_{T \wedge n}) \mathbb{1}_{\{T > n\}}]| \leq 2B \mathbb{P}[T > n] \\ &\quad \downarrow n \rightarrow \infty \\ &\quad 0 \end{aligned}$$

□